

Experimental Design and Inferential Modeling in Pharmaceutical Crystallization

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A fractional factorial experimental design was used to investigate relative effects of operating conditions on the filtration resistance of a slurry produced in a pharmaceutical semicontinuous batch crystallizer. The six operating variables were seed type, seed amount, temperature, solvent ratio, addition time, and agitation intensity. An empirical model constructed between the operating variables and filtration resistance was used to define the factory operating procedure, which reduced filtration time 3.7-fold. Several chemometric techniques were used to construct inferential models between the in-process measurement of particle chord-length distribution and filtration resistance to help detect operational problems before completing the batch and decide when batch crystallization runs should end. Depending on the model quality criterion, the most popular chemometric methods of partial least squares and top-down principal-component regression can produce lower quality models. Another chemometric approach, confidence-interval principal-component regression, predicted 70% more accurately than the best OLS model. The main effects and inferential models serve different but complementary roles in developing and implementing high-performance crystallization process operations. A main-effects model constructed from statistical experimental design data determined optimal operating conditions rapidly, while the inferential model can determine operational problems and batch end times during batch-process operations.

Introduction

Crystallization from solution is an industrially important unit operation due to its ability to provide high purity separations, which explains its predominant use in the pharmaceuticals industry. The crystal size distribution (CSD) is an important factor in the production of high-quality products and for determining the efficiency of downstream operations, such as filtration and washing. CSD-related characteristics of product crystals that have been optimized include mean crystal size (Ajinkya and Ray, 1974; Chang and Epstein, 1982; Jones, 1974), coefficient of variation (Chang and Epstein, 1982; Ma

et al., 1999), the ratio of the nucleated crystal mass to seed crystal mass (Chung et al., 1999; Mathews and Rawlings, 1998; Miller and Rawlings, 1994), and the crystal shape (Ma et al., 2000). An advantage of these objectives is their mathematical convenience, in that the objectives can be computed directly from the moments of the CSD.

The *true objective* for a pharmaceutical crystallization process may be to achieve sufficient product purity, to minimize the filtration time, or to achieve sufficient tablet stability when the crystals are mixed with crystals of other chemical species before forming the tablet. No reliable first-principles models are available for relating the CSD to these practical objec-

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tives. The development of such models is especially challenging for the crystals produced in the pharmaceutical industry, as these crystals often have a needle-like or more complex shape.

This motivates the use of statistical experimental design procedures to define optimal operating procedures for the crystallization of pharmaceutical chemicals, in which a *true process objective* is used as the goal of the optimization (Johnson et al., 1997). In this study, an unbiased multivariable screening design of experiments led to a main-effects model that was followed by process optimization to minimize the filtration resistance, which was the primary variable of importance for the industrial crystallizer of interest. A close inspection of the crystallization process for a salt of an organic molecule that had a molecular weight of less than 1000 g/mol indicated six key input variables with respect to the output variable of interest, the filtration resistance. The optimized process was verified in the laboratory. A recommended factory process was specified and successively demonstrated on the factory scale.

Another goal of this study was to construct a soft sensor (also known as an *inferential model*) to predict the filtration resistance in-process for the factory crystallizer. The purpose of the soft sensor was to aid in detecting operational problems before the completion of the batch, and in deciding when batch crystallization runs should end, when the crystals are removed from the crystallizer and filtered. Chemometric techniques were applied to data collected from laboratory experiments to construct the soft sensor, which predicts filtration resistance from a Lasentec focused-beam reflectance measurement (FBRM) instrument. This real-time in-process sensor, which obtains its information by laser backscattering, is rugged enough to be implemented on industrial crystallizers.

The challenging aspect of the model construction is that the quantity of data is low relative to the model dimensionality. In our application, the chord-length distribution was represented as a vector by binning the chord counts into 10 ranges of chord length. The number of experiments is 24. The data are highly collinear, so the model produced using ordinary least squares was too inaccurate to be useful. To address the collinearity, several chemometrics methods were used to construct inferential models, including partial least squares (PLS) and five different principal-component regression (PCR) methods. It was found that the most commonly used chemometrics methods of PLS and top-down PCR may produce lower quality predictions than those produced by alternative PCR methods. More specifically, the chemometric technique known as confidence-interval PCR gave the tightest prediction intervals.

Experimental Procedure

The pharmaceutical product was a salt of an organic molecule with molecular weight less than 1,000 g/mol. The product was crystallized in a baffled reactor, where the supersaturation was created by adding a less efficient solvent that is miscible with the original solvent. Two streams, acetonitrile (product antisolvent) and toluene product concentrate, were added simultaneously above the liquid surface at a constant rate to a seed bed of product. Each run was made

at the 40-g scale, with the final concentration between 38 and 41 g/L, for a total volume of one liter. Typical crystallizations produced a needle morphology with a length-to-width ratio greater than eight. Needle lengths reached 100 μm .

The relative effects of six input variables (operating conditions) on the main output variable, filtration resistance, were determined (Johnson et al., 1997). The input variables were the agitation intensity, seed type, seed amount, temperature, solvent ratio, and addition time. The output variables were filtration resistance (measured off-line) and particle chord-length distribution (measured in-process). Optical micrographs of the crystals were also taken.

The agitation speed was varied between 350 and 1450 rpm for a 62 mm diameter 45° pitched four-blade turbine in a 112-mm-dia. vessel. Each crystal slurry was agitated 2 h longer than the addition time of antisolvent. The seed amount (weight percentage, wt %) was based on the product in the starting concentrate. Seed was of two types: (1) final-product crystals that were dried and screened through a 20 mesh screen in a hammer mill and suspended in solvent, and (2) a stock supply of unfiltered slurry seed adjusted to the desired solvent ratio and 20 g/L starting concentration. Throughout each run, the temperature was held constant, and the concentrate and antisolvent were charged in unison to maintain a constant solvent ratio.

Measurements of filtration resistance were obtained by filtering one-half of each crystallization slurry at 138 kPa on a 50-mm-dia. filter using a tight-weave multifilament polypropylene filter medium. Determination of the average specific cake resistance and the medium resistance was based upon the traditional filtration theory (Leu, 1986). All plots of average inverse filtrate rate vs. filtrate volume were linear, giving the average specific cake resistance from the slope of the curve and the medium resistance from the intercept. For easier interpretation of experimental data, the filtration time to collect 400 mL of filtrate was calculated based on the average specific cake resistance and the medium resistance of each run. The remaining slurry was analyzed for particle chord-length distribution, which was measured *in situ* by an FBRM (model M300, version F) instrument. A chord length is a straight line between any two points on the edge of the particle or particle structure (agglomerate). The FBRM counts hundreds of thousands of chords per second. This results in a number chord-length distribution (number of counts per second sorted by chord length). The chord-length distribution is sensitive to particle shape and particle population.

The chord-length distribution can be correlated with any upstream or downstream process variable that is a function of particle shape, dimension, and/or number of particles in the particle system. The primary chord-length distribution (number of counts per second sorted by chord length into bins) can be transformed by arithmetic manipulation to a weighted distribution. The weighted distributions that were used are the length-weighted distribution, the square-weighted distribution, and the cube-weighted distribution. There is a mean chord length associated with each of these distributions.

An optical micrograph of the crystals from each slurry at 200 $\times$ magnification was taken by a Zeiss optical microscope, model number ICM 405. Results were drawn by relating the optical micrographs with filtration times and the FBRM.

Theory

Experimental design

Using statistical experimental design gives more dispersed data for correlation, that is, the experiments are selected to more adequately cover the entire range of operations. This provides more confidence when scaling up the results to the factory. Covering a wide range of operations enables the soft sensor to predict the filtration resistance more reliably for a variety of crystal slurries. It is an important point that the data used to construct a soft sensor should not only be data collected during "best practices" or around the optimal operating conditions, since a goal of the soft sensor is to provide accurate predictions even during abnormal operating conditions.

A multivariable one-eighth fractional factorial screening design of experiments was initially conducted to characterize the main effects of each input variable on the cake resistance and filtration time (Gunter, 1993). The screening design consisted of eight variable runs with two centerpoints.

It is an important point that the goal of this statistical experimental design approach is to directly optimize the filtration efficiency, whereas the goal of some other experimental design studies was to obtain improved estimates of the kinetic parameters for crystal nucleation and growth (Chung et al., 2000; Mathews and Rawlings, 1998; Miller and Rawlings, 1994).

Chemometrics

The strong correlations within the data make it impossible to construct an accurate ordinary least-square (OLS) model between the filtration time and the entire chord-length distribution. Hence ordinary least squares was only applied to relate the filtration time to certain characteristics of the chord-length distribution (such as mean chord length, fraction of small particles). The ability of the chemometric methods of PCR and PLS to handle highly correlated data allows these methods to construct inferential models based on the entire chord-length distribution.

Principal-Component Regression Methods. The first step in PCR is to transform the input variables into a set of new variables, called *principal components*, which are uncorrelated. The next step is to select as new variables a subset of the principal components that has a smaller dimensionality than the original input space. The last step is to perform ordinary least-square regression of the principal components against the output. Five different methods were considered for selecting which principal components to include in the regression model:

- Top-down selection PCR (TPCR) (Xie and Kalivas, 1997)
 - Correlation PCR (CPCR) (Xie and Kalivas, 1997)
 - Forward selection PCR 1 (FPCR1) (Draper and Smith, 1981; Phatak et al., 1993)
 - Forward selection PCR 2 (FPCR2) (Xie and Kalivas, 1997)
 - Confidence interval PCR (CIPCR)
- The inferential model has the form

$$y = x^T b, \quad (1)$$

where y is the output prediction, x is the vector of inputs, b is the vector of regression coefficients, and the data are assumed to be mean centered. The PCR method computes the regression coefficient from

$$b = V[V^T X^T X V]^{-1} V^T X^T Y, \quad (2)$$

where Y is the vector of outputs in the training set, X is the matrix of inputs in the training set, and the columns of V are the eigenvectors of $X^T X$ corresponding to the selected principal components. If all principal components were used, then the estimate would be the same as in ordinary least squares. Only a subset of the principal components can be used for regression due to high data collinearity and the small quantity of data relative to the number of regression coefficients.

The challenge then becomes to select the best principal components for use in the regression step. Top-down selection, the most popular method for principal-component selection, assumes that the most informative principal components are the ones with the greatest variability, and then uses some criterion for deciding the number of principal components. However, the principal components with the highest variability may not be the most informative for the purpose of prediction, since this criterion does not contain any information on the correlation between each principal component and the output (Jackson, 1991; Jolliffe, 1982; Massy, 1965). In order to choose which of these principal components represent better the correlation of the input to the output space, CPCR (Xie and Kalivas, 1997), FPCR1 (Draper and Smith, 1981; Phatak et al., 1993), FPCR2 (Xie and Kalivas, 1997), and CIPCR have been proposed.

Since the other chemometrics methods have clear descriptions in other journal articles, only the CIPCR method will be described here. The CIPCR method is designed specifically to give tight prediction intervals. CIPCR constructs confidence intervals for regression coefficients between the principal components and the output, and selects only those that are statistically significant. The significance is determined using a hypothesis test based on the variance of the regression coefficients computed using standard statistical formula (Beck and Arnold, 1977).

Partial Least-Square Regression. The PLS method uses output data Y while it is constructing the new regression variables and thus eliminates the weight of variables with high variability but no relevance to the output space (Martens and Naes, 1989; Russell et al., 2000). Due to the nonlinear form of the PLS estimator, it is difficult to derive exact prediction intervals. An approximation on the prediction intervals can be constructed by linearization of the PLS estimator (Phatak et al., 1993). The accuracy of the prediction intervals was comparable to Monte Carlo simulation for several data sets (Phatak et al., 1993).

Measures of Accuracy of Models. The predictive ability of the model can be quantified by the relative ability of prediction (RAP) (Martens and Naes, 1989). The RAP is close to unity for perfect predictors, and is close to zero for very poor predictors. The mean value of the prediction interval is another property of interest that is used as a criterion to compare between the different methods (such as Faber and

Kowalski, 1996; Jackson, 1991; Phatak et al., 1993, and citations therein). All chemometric calculations were performed using home-grown MATLAB code, except for the PLS algorithm, in which the Matlab PLS Toolbox 2.0 was used (Wise and Gallagher, 1998).

Results

Main effects model

Based on the ten experiments (described in the Experimental Design section), an ordinary least-square model (often called a *main effects* model) was fit to correlate the filtration time to the six input variables. The model indicated that in order to reduce filtration time, the agitation speed and solvent ratio should be lowered, the temperature, seed amount, and addition time should be increased, and dried seed should be preferred over a slurry seed.

It was attempted to reduce filtration time by adjusting each input variable in the direction dictated by the main effects model of the screening design. Only the addition time was not optimized to shorten run times due to project time limitations.

The first optimization run, OPT1 in Table 1, indeed produced the second best filtration time of all the runs, 8.7 min. In order to verify the predictions of the main effects model, additional runs were made, changing one input variable of the optimized run, OPT1, to the negative of the design range. In all cases, except for seed type, switching only one variable to the negative of the optimum predicted by the screening resulted in an increase in filtration time. In addition, multiple inputs were changed to the negative of the optimum settings, run OPTNEG of Table 1, resulting in the third longest filtration time of all the runs, 99 min.

An empirical model to predict the filtration time of a dried seed run over the ranges evaluated in Table 1 was developed using first-order multivariable ordinary least-square regression. The model was based on the 16 crystallization runs, including the optimization runs. The slurry seed runs were excluded from the model to increase the model accuracy, as will be discussed below.

The prediction equation of filtration resistance from the five input variables (excluding seed type) is given by:

$$\begin{aligned} \text{Time (min) to collect} &= 34.5 + 0.0362 \cdot (\text{agitation intensity, rpm}) \\ &+ 21.47 \cdot (\text{solvent ratio}) - 1.77 \cdot (\text{temp., } ^\circ\text{C}) - 2.64 \cdot (\text{charge time, h}) \\ &- 0.722 \cdot (\text{seed amount, wt. \%}) \end{aligned} \quad (3)$$

With respect to the ranges studied, agitation intensity had the largest effect on the resulting filtration time of a crystallization, followed by solvent ratio, temperature, charge time, and seed amount.

A strong interaction between seed type and seed amount exists. The amount of seed had little effect when the dried seed was used, but if slurry seed was used, the amount of seed had a dramatic effect. The adverse combination of low seed amount and slurry seed led to the longest filtration times of all the runs, 191 min. Conversely, the best filtration time, 6.0 min, of all runs was experienced using 10 wt % slurry seed under favorable conditions for all the other variables, except addition time, similar to run OPT1 in Table 1. It was our hypothesis that when there is not enough slurry seed, the crystallization system becomes starved for crystal surface area, resulting in a high supersaturation and significant nucleation. According to this hypothesis, we would expect to see a bimodal distribution for crystals produced from crystallization runs with low slurry seed (this is confirmed in optical micrographs below). The dried seed was more desirable, due to the strong sensitivity of filtration time to the amount of slurry seed.

The difference in sensitivities for the dried and slurry seed is likely due to the different nature of the crystal surfaces. The surface of the dried seed is rough and fractured during the drying and milling process, which undoubtedly results in a relatively large number of growth sites distributed across the surface even if the amount of dried seed is low. On the other hand, the slurry seed is quite intact, with a much smoother surface. Since the crystals grow in a needle-like shape, the growth sites for slurry seed are predominantly at the opposite ends of the longest length dimension of the crystal. When the quantity of slurry seed is low, the limited num-

Table 1. Experimental Design

	Screening Design Range	OPT1	OPTNEG	Previous Factory	Recommended
<i>Input variables</i>					
Agitation intensity, rpm	350 to 1450	350	1450	1450 (tip speed)	900
Seed amount, wt. % of batch	3.0 to 10	10	3	5.3	10
Seed type	Slurry or dried	Dried	Dried	Dried	Dried
Temperature, °C	15 to 25	25	16	18	24
Volume ratio, MeCN:Tol	1:1 to 2.33:1	1:1	2.33:1	1.66:1	1.22:1
Charge time, h	3 to 9	3	3	6.9	6.9
<i>Output variables</i>					
Avg. specific cake resistance, $\times 10^{-10}$, m/kg	290 to 8.5	12	150	89	35
Medium cake resistance, $\times 10^{-10}$, 1/m	62 to 3.2	7.9	36	12	11
Filtration time,* min.					
Experimental result	191 to 6.0	8.7	99	57	23
Dried model prediction		8.8	98	68	25

* Filtration time for collecting 400 mL of filtrate.



Figure 1. Optical micrograph (200 $\times$); slowest filtration.

ber of growth sites results in an increase in supersaturation during the crystallization run, resulting in nucleation and the bimodal distribution of crystals.

Optical micrographs

Optical micrographs confirmed an obvious bimodal distribution of particles and fines existing in the two runs with the highest filtration resistance, which used slurry seed (see Fig-

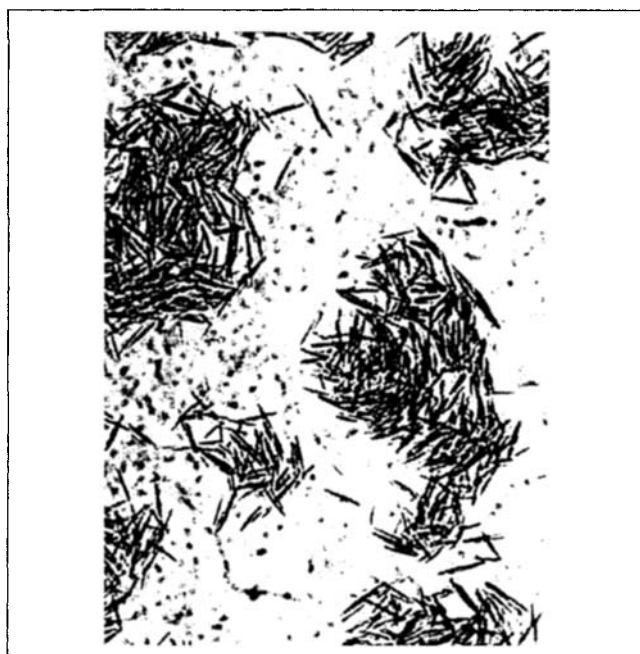


Figure 2. Optical micrograph (200 $\times$); laboratory recommended.



Figure 3. Optical micrograph (200 $\times$); factory recommended.

ure 1). Conversely, optical micrographs confirmed that the lowest filtration resistance corresponded to cases where few fines existed and particles were of considerable size, 50–100 μm in length (Figures 2–4). Comparing the results in Table 2 with the optical micrographs, we can qualitatively correlate the properties of filtration time, the percentage of particles with chord length between 0.8 and 1.9 μm , and mean chord length with the crystal size distribution in the related images.



Figure 4. Optical micrograph (200 $\times$); fastest filtration.

Table 2. Filtration Time and Chord Length Distribution Characteristics

	Slowest Filtration	Laboratory Recommended	Factory Recommended	Fastest Filtration
Filtration time,* min	191	26	12.7	6
% particles with $0.8 \leq \text{chord length} \leq 1.9 \mu\text{m}$	9.2	8.2	6.7	6.5
Mean chord length of primary CLD	8.9	11.0	11.1	13.0

* Filtration time for collecting 400 mL of filtrate.

Higher filtration times are associated with smaller crystals and the presence of fines.

Factory Procedure and Scale-Up. In order to maintain compliance with factory processing constraints for the product, a recommended factory procedure was developed, shown in Table 1. In the laboratory, the recommended conditions reduced the filtration time 2.5-fold, from 57 to 23 min.

The recommended factory procedure was demonstrated successfully four times at the 240-kg scale, which is 6000-fold greater than the laboratory scale. Samples of the slurry were taken and filtered in the laboratory, giving filtration resistance equivalent to 12–22 min, with an average of 16 min to collect 400 mL of filtrate. This result was better than the 23 min expected based on three 40-g-scale runs.

The previous factory procedure was run side by side with the recommended factory procedure in the laboratory, yielding an improvement from 57 min to 23 min (Table 1). Laboratory filtration analysis of the previous factory procedure at the factory scale does not exist, but factory perforated-basket centrifugation times for the product, not including downtime, were reduced from 110 h for the previous factory procedure to 30 h by implementing the recommended factory procedure, a 3.7-fold reduction in filtration time.

Inferential Modeling. The primary goal of the main-effects model was to define optimal operating conditions for the semicontinuous batch-crystallization process. One result from the study was the empirical model (Eq. 3) relating the filtration time to five key variables that defined the operating procedure. The goal of the inferential model, also known as a soft sensor, is to relate the filtration time to the chord-length distribution, which can be measured during the crystallization run. Such an inferential model can be used to alert process operators of operations problem before the completion of the batch, and to decide when crystallization runs should end, when the crystals are removed from the crystallizer and filtered. The main effects and inferential models serve different but complementary roles in the development and implementation of high-performance crystallization process operations.

Instead of filtration time, better results were obtained when the soft sensors were constructed to predict the logarithm of the average specific cake resistance. The entire data set of 24 experiments was divided into calibration and validation sets using eight different combinations, thus creating eight distinct subdivisions of the experimental data into calibration and validation sets. The eight distinct subdivisions were created in order to assess the sensitivity of the inferential models on the particular calibration set used. Methods that consistently give high-quality inferential models independent of the particular calibration-validation set are preferred, as their results are more reliable.

To serve as a benchmark, inferential models were constructed using the ordinary least-square method. Four OLS models were constructed for which the predicted variables were (1) the percentage of particles with a chord length between 0.8 and $1.9 \mu\text{m}$ and the mean particle chord length associated with the (2) primary distribution, (3) length-weighted distribution, and (4) square-weighted distribution.

The relative ability of prediction and mean range of prediction intervals (confidence level 95%) for the OLS models are given in Tables 3 and 4. Of the four predictor variables, the mean chord length of the primary distribution and the percentage of particles with chord length between 0.8 and $1.9 \mu\text{m}$ give higher quality models. In all cases, the relative ability of prediction is low ($\text{RAP} \leq 0.8$) and the prediction intervals are relatively wide, which limits the usefulness of the OLS models. Also, the quality of all OLS models depends strongly on the selection of the calibration-validation set. Of the predictors used in the OLS analysis, the most reliable predictors are the percentage of small particles and the mean chord length of the primary chord-length distribution. The trends are consistent with those observed from the optical micrographs.

Inferential models were constructed between the chord-length distribution and the average specific cake resistance using the PLS and PCR methods. The variance of the measurement error, which is needed by some of the PCR methods, was estimated from the centerpoint runs in the frac-

Table 3. Relative Ability of Prediction for OLS Models

Predictor Variable	Calibration-Validation Subdivision								Avg.
	1	2	3	4	5	6	7	8	
% particles with $0.8 \leq \text{chord length} \leq 1.9 \mu\text{m}$	0.4	0.7	0.8	0.7	0.7	0.8	0.7	0.8	0.7
MCL of primary CLD	0.6	0.8	0.8	0.7	0.2	0.8	0.4	0.8	0.6
MCL of length-weighted CLD	0.6	0.7	0.8	0.6	-0.1	0.6	0.2	0.7	0.5
MCL of square-weighted CLD	-0.4	0.1	-0.8	0.1	-0.7	-0.4	-0.5	-1.1	-0.5

MCL is the mean chord length, and CLD is the chord-length distribution.

Table 4. Mean Range of Prediction Intervals (Confidence Level 95%) for OLS Models

Predictor Variable	Calibration-Validation Subdivision								Avg.
	1	2	3	4	5	6	7	8	
% Particles with $0.8 \leq \text{chord length} \leq 1.9 \mu\text{m}$	2.1	2.2	1.9	2.4	2.2	2.0	2.3	2.1	2.2
MCL of primary CLD	2.1	2.3	2.3	2.0	2.0	1.8	1.9	2.1	2.1
MCL of length-weighted CLD	2.5	2.5	3.1	2.1	2.2	2.3	2.2	2.7	2.5
MCL of square-weighted CLD	3.9	3.6	3.9	3.7	3.2	3.4	3.5	3.4	3.6

MCL is the mean chord length, and CLD is the chord-length distribution.

Table 5. Number of Predictor Variables

	Calibration-Validation Subdivision								Avg.
	1	2	3	4	5	6	7	8	
TPCR	8	4	3	3	3	1	6	3	3.9
CPCR	3	2	3	1	3	2	4	4	2.8
FPCR1	4	2	1	3	2	3	3	2	2.5
FPCR2	4	4	4	2	5	3	3	4	3.6
CIPCR	4	3	1	3	2	3	6	2	3.0
PLS	5	4	1	2	6	1	5	3	3.4

tional factorial design of experiments. The data (particle chord-length distribution) were mean-centered before the construction of the predictors. The complete data set included 24 experiments, with 10 bins in the chord-length distribution (number of predictor variables at the original space). The rank of the covariance matrix $X^T X$ was less than 10. The strong collinearity of the data and the relatively small number of experiments pose a challenging problem for constructing a reliable and accurate predictor.

The number of predictor variables used in each chemometric technique is reported in Table 5. On average, TPCR selects a large number of principal components than the other methods. The principal components selected as predictor variables in each PCR method are reported in Table 6. In almost all cases the alternative PCR methods do not include the second or fourth principal components in constructing the inferential models. The implication is that the second and fourth principal components are not strongly correlated to the predicted output (logarithm of the filtration resistance). In fact, the eighth principal component, which has one of the smallest variances, is selected much more often than the second principal component, which has one of the largest variances. The TPCR method has on average a larger number of principal components because its ordering of the principal components by variance results in the inclusion of the second

and fourth principal components, which do not add significantly to the quality of the inferential model. Other examples from chemical engineering, meteorology, and economics are available where principal components of smaller variance can be more important for prediction than principal components of larger variance (Jolliffe, 1982; Massy, 1965). None of these examples involve data that could be considered bizarre or obscure, rather, it is conjectured that such data sets are common in practice (Jolliffe, 1982).

The regression coefficient b_1 is consistently positive for all of the data subdivisions, indicating that the filtration resistance increases when the number of small crystals increases, which agrees with the optical micrographs and industrial experience. There were no other consistent trends among the rest of the regression coefficients.

Except for TPCR, the models constructed using the chemometric methods gave higher average relative ability of prediction than the OLS models, and the RAP was also less sensitive to the selection of the calibration-validation set (see Table 7). The worst RAP among the inferential models was found for PLS and TPCR, which are the most popular chemometrics methods. The best RAP was recorded for FPCR2, which was greater than 0.7 for all data subdivisions. The high RAP for FPCR2 can be attributed to the fact that it uses the validation set during its selection of principal components. A weakness of FPCR2 is that it gives wider prediction intervals than some other PCR methods (see Table 8).

The PCR models had tighter prediction intervals than for the models constructed using OLS and PLS. PLS gave much larger prediction intervals than the other chemometric methods. The PCR method proposed in this article, CIPCR, gave the tightest prediction intervals. This is not surprising, as the ability of a principal component to predict the output is used to select the principal components in CIPCR. Figure 5 shows the measured value, estimated value, and prediction intervals for CIPCR applied to the validation set for data subdivision 1.

Table 6. Selected Principal Components, Where "1" Refers to the Principal Component with Largest Variance, "2" Refers to the Principal Component with Second Largest Variance, and So On

	Calibration-Validation Subdivision							
	1	2	3	4	5	6	7	8
TPCR	1-8	1-4	1-3	1-3	1-3	1	1-6	1-3
CPCR	1,3,5	1,3	1,3,5	1	1,3,8	1,3	1,3,5,8	1,3,6,7
FPCR1	1,3,5,4	1,3	1	1-3	1,3	1,3,8	1,3,5	1,3
FPCR2	1,3,5,8	1,3,5,2	1,3,5,6	1,3	1,3,8,6,2	1,3,6	1,3,8	1,3,6,7
CIPCR	1,3,4,5	1,3,8	1	1-3	1,3	1,3,8	1,2,3,5,6,8	1,3

Table 7. Relative Ability of Prediction for Chemometric-Methods

	Calibration-Validation Subdivision								Average
	1	2	3	4	5	6	7	8	
TPCR	0.9	0.6	0.9	0.7	0.8	0.7	0.7	0.9	0.7
CPCR	0.9	0.9	0.9	0.6	0.9	0.9	0.8	0.9	0.8
FPCR1	0.9	0.9	0.7	0.7	0.7	0.7	0.8	0.9	0.8
FPCR2	0.9	1.0	0.9	0.8	1.0	0.9	0.9	0.9	0.9
CIPCR	0.9	0.8	0.7	0.7	0.9	0.7	0.8	0.9	0.8
PLS	0.9	0.6	0.7	0.6	0.9	0.7	0.7	0.9	0.8

The performance of all of the methods showed some sensitivity to the selection of the calibration and validation sets. This is because the number of experimental samples was small compared to the number of predictor variables, and because of the high data collinearity. The subdivisions were used here to compare and evaluate various chemometric methods in order to select the best technique for this particular application. The final inferential model should be constructed using CIPCR applied to all of the crystallization data.

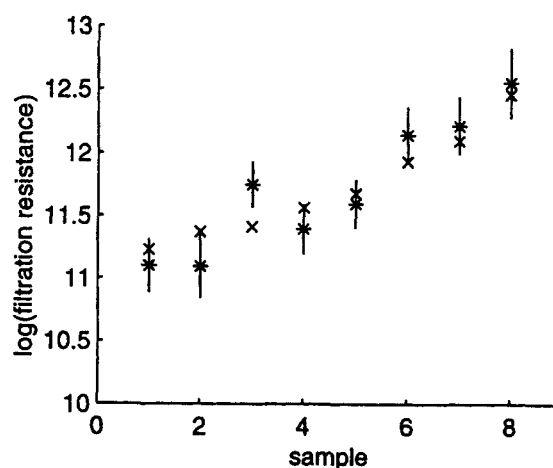
Conclusions and Recommendations

Using an unbiased multivariable testing fractional factorial experimental design, where many variables were changed simultaneously, allowed rapid optimization over six variables for a semicontinuous batch crystallization process. An empirical model was developed to relate five of the input variables to filtration time. A recommended operating procedure was determined by optimizing the empirical model over the input variables. Scale-up of the recommended process by 6000-fold was successful, resulting in a 3.7-fold (73%) reduction in perforated-basket centrifugation time compared to previous factory batches.

The recommendations obtained by the empirical model in order to decrease filtration resistance were consistent with crystallization theory. Theory dictates enhanced growth conditions and reduced nucleation to create a more monodisperse size of particles and fewer fine particles. A large number of seed and growth sites were required to reduce nucleation, which reduced bimodal particle formation and reduced filtration resistance. Within the ranges studied, dried seed was more robust and preferred over slurry seed. For the dried seed, agitation intensity was the most important variable affecting the resulting filtration time of a crystallization slurry. Low agitation intensity was recommended, since it limits secondary nucleation and promotes a monodisperse particle size.

Table 8. Mean Range of Prediction Intervals (Confidence Level 95%) for Chemometric Methods

	Calibration-Validation Subdivision								Average
	1	2	3	4	5	6	7	8	
TPCR	0.55	0.71	0.70	0.56	0.63	0.80	0.49	0.69	0.64
CPCR	0.45	0.66	0.76	1.05	0.59	0.72	0.48	0.79	0.69
FPCR1	0.43	0.66	0.71	0.56	0.61	0.26	0.51	0.68	0.60
FPCR2	0.49	0.74	0.95	0.94	0.61	0.85	0.62	0.79	0.75
CIPCR	0.43	0.62	0.71	0.56	0.61	0.62	0.39	0.68	0.58
PLS	0.96	1.23	0.94	0.81	1.39	0.98	0.88	1.46	1.08

**Figure 5. Measured output (x), estimated output (*), and prediction intervals for log (average specific cake resistance) vs. sample number.**

Both reducing the solvent ratio (more toluene) and increasing batch temperature led to increased solubility and so reduced supersaturation of the product, which favors the reduction of filtration time. Extension of the addition time also promoted controlled growth by reducing supersaturation.

The utility of chemometrics for the in-process prediction of crystal properties was demonstrated through application to the crystallization of a pharmaceutical. The OLS, PLS, and several PCR methods were compared in terms of the accuracy of their predictions. The chemometric methods that used multiple bins from a chord-length distribution gave a higher and more reliable relative ability of prediction than the OLS methods developed based on a unique predictor variable. The most popular chemometrics techniques, PLS and TPCR, did not give the most accurate predictions. CIPCR gave the smallest prediction intervals, with its predictions being 70% more accurate than the predictions of the best OLS model.

The use of rational experimental design combined with enhanced chemometrics techniques allows the engineer to more accurately predict and control challenging pharmaceutical crystallization processes. The main effects and inferential models serve different but complementary roles in the development and implementation of high-performance crystallization process operations. A main-effects model constructed from data collected from statistical experimental designs allows the rapid determination of the optimal operating conditions. The inferential model allows the rapid determination of operational problems and end times during batch-process operations.

Although CIPCR outperformed TPCR and PLS for these particular data sets, this will not necessarily be true for other data sets. However, the results do suggest a general procedure for constructing the most accurate and reliable inferential models. Several chemometrics methods should be applied to a given data set, to result in multiple inferential models. These chemometrics methods should include not only the popular PLS and TPCR methods, but also include alternative PCR methods that select principal components based on the quality of the inferential model, rather than just on variability

of the input space. The best inferential model should be selected based on optimizing some well-defined model quality criterion, such as the average size of the prediction intervals.

Acknowledgments

Funding was provided by the Merck Foundation. Rich Becker of Lasentec is acknowledged for some helpful comments, especially with regard to the FBRM instrument.

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Manuscript received Jan. 24, 2000, and revision received June 26, 2000.